



The Role of Social Media Use and Associated Risk and Protective Behaviors on Depression in Youth Adults: A Longitudinal and Network Perspective

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Accepted: 22 April 2024

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Abstract

Social media use has been associated with depression in youth. It remains unclear how individual differences in social media use and depression relate to other health-related behaviors such as physical activity, green space exposure, cannabis use, and eveningness. We aimed to examine these associations and network structure with the Green/Screen Study, a Canadian cohort of 425 adults (82.5% female) aged 18–25 years who completed three online questionnaires between May 2021 and January 2022. Longitudinal analyses showed sex differences in the association of social media use with depression. Although social media use was positively associated with depression after adjusting for other health-related behaviors, it did not predict an increase in depressive symptoms over time. Shorter time intervals and more detailed measures of social media use may be needed to capture social media context- and usage-specific associations with depression.

Keywords Social media · Depression · Cannabis use · Green space exposure · Physical activity · Young adults

Over the past two decades, there has been a significant increase in the prevalence of depression in adolescents and young adults (Keyes & Platt, 2023; Yu et al., 2020). Simultaneously, there has been an unprecedented inclusion of technology in our daily social lives. Digital media use among adults in the United States (USA) has increased from 5 to 72% in only two decades (Pew Research Center, 2021) and newer generations of youth have

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experienced even higher rates of digital media use, with 95% of the adolescent population using digital media (Anderson & Jiang, 2018).

Social media is a subset of digital media that has become increasingly popular over the past two decades. The rates of social media use in the U.S. are 72% in the general population and 84% among youth aged 18 to 29 years, with similar rates in other countries, including Canada (Wike et al., 2022). There has been a recent increase in social media use during and post-COVID-19 pandemic, likely related to social isolation, boredom, and a desire for social connectedness (Avena et al., 2021). While causation has not been established, some researchers argue that the increased depression rates are at least partly attributable to the expansion of smartphone use among youth (Twenge & Campbell, 2018; Twenge et al., 2018). In the U.S., female-specific increases in rates of suicide (a possible outcome of severe depression), by 87% in girls aged 15–24 (Garnett et al., 2022) have mirrored increases in the use of social media platforms (Pew Research Center, 2023; Twenge et al., 2018). Many mechanisms have been hypothesized as potentially influencing the associations between social media use and depression, including cyberbullying, social contagion, increased opportunities for upward social comparisons (Vidal et al., 2020), and sleep disruption (Alonzo et al., 2021). Some studies in youth support a longitudinal association between social media use and depression and other emotional and behavioral problems (Boers et al., 2019; Riehm et al., 2019), especially when the time spent on social media is excessive (McNamee et al., 2021). Yet, other studies show no (Beeres et al., 2021; Coyne et al., 2020; Orben & Przybylski, 2019) or weak longitudinal associations (Heffer et al., 2019; Puukko et al., 2020).

Young adulthood is a key developmental period in the shaping of behaviors that can affect health later in life (Hawkins et al., 1992). One of the concerns with excessive use of social media is the displacement of other health-related behaviors, known to have positive effects on mental health, such as time outdoors and exposure to green spaces (Liu et al., 2023), sleep (Scott et al., 2021), and exercise (Pearce et al., 2022). Displacement theory postulates that social media use may directly have an effect on time spent away from physical activities, in-person socializing, or other health-promoting activities such as sleep and schoolwork, with potential negative impacts on the person's mental health (Hall & Liu, 2022). Conversely, maintaining a healthy lifestyle can positively affect mental health (Maenhout et al., 2020).

While often studied separately, health-related behaviors are interrelated and associated with mental health outcomes. For example, eveningness, or the tendency to sleep late in the night and wake up late in the morning, has been associated with greater use of digital media, greater sedentary behaviors, and more depressive symptoms (Au & Reece, 2017; Patterson et al., 2016). The increase in social media use during the pandemic was associated with worse sleep patterns in students (Ali et al., 2022), and sleep contributes to the relationship between sedentary behaviors and mental health symptoms (Werneck et al., 2020). Situations leading to stress and isolation can further exacerbate behavioral and substance use disorders (Avena et al., 2021). During the COVID-19 pandemic, youth displayed an increase in sedentary behaviors, reduced physical activity, disrupted sleep patterns, and a change in dietary habits (Avena et al., 2021; Pérez-Carbonell et al., 2020); these changes were paralleled by increases in screen use and mental distress, either linked to these changes in habits, or due to other causes (Caroppo et al., 2021). Furthermore, environmental features, such as exposure to green spaces, may also contribute to depression. For example, a recent meta-analysis found lower risk of depression with a 10% increase in the proportion of residential green space (Liu et al., 2023). Finally, changes in cannabis regulation have been postulated to contribute to the increase in suicide rates among

youth (Hammond et al., 2023), and there are well-known associations between cannabis and depression (Arias et al., 2020; Bolanis et al., 2020; Gobbi et al., 2019). Together, these interrelated behaviors may underlie some of the associations between social media use and depression.

Despite the growth in the number of longitudinal studies exploring the associations between social media use and depressive symptoms, the directionality and strength of association in the context of other risk and protective health-related behaviors need further exploration. The aim of this study was to examine the longitudinal and network associations between social media use and depression in youth after accounting for other health-related behaviors and sex differences.

Hypothesis

We hypothesized that higher initial levels of social media use would be associated with higher levels of depressive symptoms over time, taking into consideration baseline levels of physical activity, green space exposure, cannabis use and eveningness, with a stronger association in female participants. To better understand how health-related behaviors relate to each other and with depressive symptoms, we further explored their interrelationships using network analyses.

Method

Participants

Participants were from the Green/Screen Study (<https://osf.io/y2b9z/>), a cohort of 425 (82.5% female) young adults (18–25 years) residing in the Canadian province of Québec and speaking one of the two official languages of the country, French or English.

Participants were recruited through advertisements targeting the aforementioned age range and geographical region through four popular social media platforms (Facebook, Instagram, Twitter, and Reddit) and two classified advertisement websites (Kijiji and Craigslist). Individuals were oriented to the study website for project information and consent. All participants provided their written consent electronically. Draws of 20 lots of \$25 CAD after the first study time point and 10 lots of \$100 CAD after the third time point were offered to increase participation and reduce attrition. The study was approved by the institutional review board of the University of Quebec at Montreal (#4552_e_2021).

Measures

Participants were assessed over three time points during the following periods: May 27 to June 14, 2021 (T1), September 8 to October 13, 2021 (T2), and December 2 to January 11, 2022 (T3). Questionnaires were administered online through a custom survey platform (<https://www.elaborer.org>). At each time point, self-reported measures included depressive symptoms, social media use, greenspace exposure, physical activity, and cannabis use. Age, sex (male or female), educational attainment, and eveningness chronotype were measured at baseline.

Depressive symptoms were measured at each time point with the 9-item Patient Health Questionnaire (PHQ-9) (Kroenke et al., 2001). The PHQ-9 consists of 9 questions related to symptoms of depression on a scale from 0 (not at all) to 3 (nearly every day), with a total score ranging from 0 to 27.

Social media use was measured using the following item adapted from the Coronavirus Health and Impact Survey (Nikolaidis et al., 2021): “During the past 3 months, how much time per day did you spend using social media (such as Facetime, Facebook, Instagram, Snapchat, Twitter, TikTok)?”. Response options were “never/did not use”, “under 1 h”, “1–3 h”, “4–6 h”, and “more than 6 h.” To capture unequal distances between these response categories, they were re-coded as numeric values according to their midpoints: 0, 0.5, 2, 5, and 7 h per day, respectively.

Greenspace exposure was measured using two items adapted from Coldwell & Evans (Coldwell & Evans, 2018): “Over the last 3 months, how often did you visit outdoor spaces in the countryside (such as a forest, river or field)?” and “[...] how often did you visit outdoor green spaces in your city (such as a park, garden or canal)?”. For each item, response options were “Never”, “Once or twice”, “Every month”, “Every 2 weeks”, “Every week”, “Several times a week”, and “Daily”. These categories were re-coded numerically to indicate the average number of days of exposure per week. The two item scores (rural and urban green space exposure) were summed to obtain the total frequency of green space exposure (range, 0–7 days/week).

Physical activity was evaluated using the International Physical Activity Questionnaire – Short Form (Craig et al., 2003). This questionnaire includes seven items measuring weekly time spent on low, moderate, and high-intensity physical activity. Responses are aggregated to estimate the total metabolic equivalent task (MET)—minutes/week.

Frequency of cannabis use in the past 2 weeks was reported on a 4-point scale (“never”, “1–2 times per week”, “3 or more times per week”, and “everyday”) (Landry et al., 2004). Due to skewed responses, cannabis use was dichotomized as any versus none in regression models.

Eveningness chronotype was measured by assessing typical sleep habits using the Munich Chronotype Questionnaire (Roenneberg et al., 2003). As recommended by the questionnaire creators (Roenneberg et al., 2004), we calculated chronotype as the midsleep point on days without engagements (i.e., without work or school) and corrected for sleep debt accumulated during days with engagements. The midsleep point corresponds to the time of the day between the onset and end of the sleep period (range, 0–24 h). A higher value indicates greater eveningness, defined as the tendency to initiate sleep later in the night and to wake up later in the morning.

Statistical Analysis

Analyses were conducted in R version 4.1.2 (R Core Team, 2021). Descriptive analyses were performed to report characteristics of the sample according to attrition status. To evaluate associations between social media use at T1 and depressive symptoms at T1, T2, and T3, we used linear mixed models with random intercepts. All models were adjusted for time (T1, T2, or T3), age, and sex. We subsequently included the other health-related behaviors at T1 (physical activity, greenspace exposure, cannabis use, and eveningness) as covariates. We explored sex differences by testing the interaction between sex and social media use. To evaluate whether higher social media use at T1 was associated with a trajectory of increasing depressive symptoms over time, we tested the interaction of time with

social media use at T1. We used the same approach to evaluate whether depressive symptoms at T1 were associated with social media use over time. Confidence intervals of 95% not overlapping the null and p -values < 0.05 were considered statistically significant.

To examine the network structure of associations between depressive symptoms, social media use, and other health-related behaviors, we generated Gaussian graphical models using the qgraph (Epskamp et al., 2012) and bootnet packages (Epskamp et al., 2018). In network models, variables are visually represented as nodes, and the connections between variables are represented as edges. Edges indicate partial correlations between any two nodes after conditioning on all the other nodes in the network. To estimate edge weights, we applied the ggmModSelect algorithm (stepwise estimation) given its good performance at low sample sizes (Isvoranu & Epskamp, 2023), and we used the Spearman correlations to accommodate non-parametric associations. With ggmModSelect, the inclusion or removal of edges from the network is based on minimizing the extended Bayesian information criterion. We generated separate cross-sectional models for data at T1, T2, and T3, restricting the analyses to participants with complete data on network variables at each time point. Because chronotype is thought to be stable over time and was only measured at T1, we used the same measure for subsequent time points. Non-parametric bootstrapping with 1000 iterations was performed to evaluate the accuracy of network estimates at T1. Edge weights were considered significant if they were not set to zero in $> 50\%$ of non-parametric bootstrapped models. Bootstrap analyses were not performed on networks at T2 and T3 given smaller sample sizes. Codes are available at <https://osf.io/y2b9z/>.

Results

Sample Characteristics

Sample characteristics are presented in Tables 1 and S1. The analytic sample included 376 participants who had complete data on all measures at T1. Of this sample, 233 (62.0%) completed at least one of the follow-up assessments (T2 or T3). These participants, compared with those who only completed T1 assessments, were more likely to be female and less likely to use cannabis and reported fewer depressive symptoms.

Longitudinal Depressive Symptoms as a Function of Social Media Use at T1

Greater social media use at T1 was associated with more depressive symptoms on average over T1, T2, and T3 (Table 2). The association remained significant after adjusting for other health-related behaviors at T1. Cannabis use and greater eveningness were also associated with more depressive symptoms. Sex, greenspace exposure, and physical activity were not significantly associated with depressive symptoms. Depressive symptoms tended to increase over time, but the slope of time was not statistically significant.

The interaction between social media use and time was not significant: coefficient = -0.86 (95% CI, $-0.244, 0.071$), indicating that greater social media use at T1 was not associated with an increase or decrease in depressive symptoms over time (Fig. 1). There was however a significant interaction of social media use with sex: coefficient = -0.785 (95% CI, $-1.434, -0.136$). In female participants, social media use at T1 was significantly associated with more depressive symptoms on average: coefficient = 0.551 (95% CI, $0.264, 0.838$). The coefficient indicates that for each additional hour

Table 1 Description of the sample at baseline and characteristics associated with follow-up participation

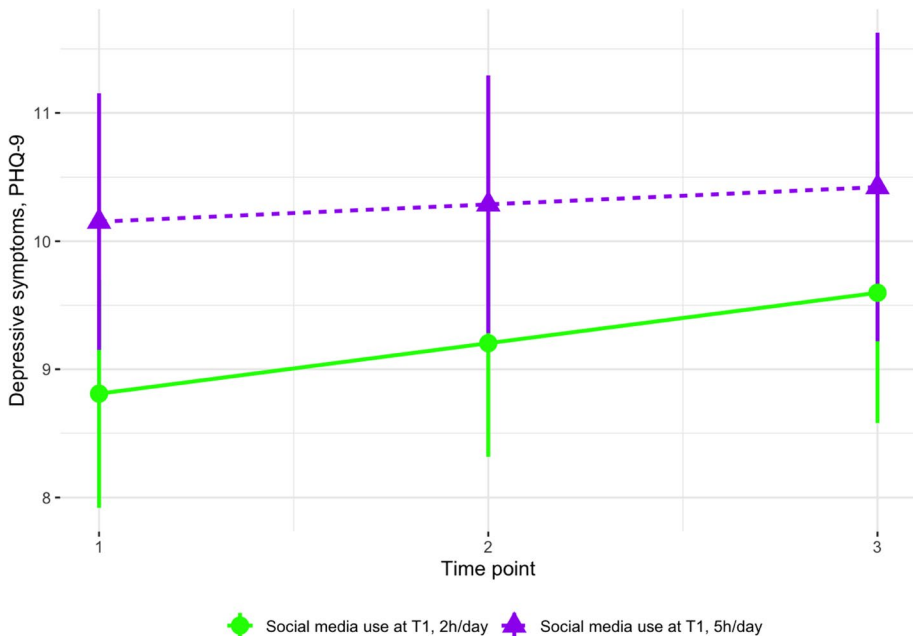
	Characteristics of the baseline sample (N = 376)		Completed T2 and/or T3?		p-value
	Yes (N = 233)	No (N = 143)	Yes (N = 233)	No (N = 143)	
Age in years, median [25th;75th]	22.0 [20.0;24.0]	22.0 [20.0;24.0]	22.0 [19.0;24.0]	22.0 [20.0;24.0]	0.865
Sex, N (%):					0.019
Female	310 (82.4%)	310 (82.4%)	201 (86.3%)	109 (76.2%)	
Male	66 (17.6%)	66 (17.6%)	32 (13.7%)	34 (23.8%)	
Educational attainment, N (%):					0.053
High school or lower	99 (26.4%)	99 (26.4%)	53 (22.7%)	46 (32.4%)	
Some college or higher	276 (73.6%)	276 (73.6%)	180 (77.3%)	96 (67.6%)	
Employment, N (%):					0.823
None	109 (29.1%)	109 (29.1%)	68 (29.4%)	41 (28.7%)	
Half time	176 (47.1%)	176 (47.1%)	106 (45.9%)	70 (49.0%)	
Full time	89 (23.8%)	89 (23.8%)	57 (24.7%)	32 (22.4%)	
Student status, N (%):					0.367
None	80 (21.6%)	80 (21.6%)	44 (19.2%)	36 (25.4%)	
Half time	43 (11.6%)	43 (11.6%)	28 (12.2%)	15 (10.6%)	
Full time	248 (66.8%)	248 (66.8%)	157 (68.6%)	91 (64.1%)	
Depressive symptoms, median [25th;75th]	8.00 [4.00;13.0]	8.00 [4.00;13.0]	7.00 [4.00;12.0]	9.00 [5.00;14.0]	0.034
Social media use (hours/day), median [25th;75th]	2.00 [2.00;5.00]	2.00 [2.00;5.00]	2.00 [2.00;5.00]	2.00 [2.00;5.00]	0.642
Physical activity (MET minutes/week), median [25th;75th]	1668 [792;3388]	1668 [792;3388]	1508 [808;3252]	1822 [777;3636]	0.349
Green space exposure (days/week), median [25th;75th]	1.12 [0.50;3.73]	1.12 [0.50;3.73]	1.12 [0.50;3.62]	1.23 [0.62;3.73]	0.224
Chronotype (midpoint of sleep), median [25th;75th]	4.49 [3.62;5.64]	4.49 [3.62;5.64]	4.43 [3.71;5.38]	4.51 [3.54;6.02]	0.607
Cannabis use (once or more/week), N (%):					0.001
No	284 (75.5%)	284 (75.5%)	190 (81.5%)	94 (65.7%)	
Yes	92 (24.5%)	92 (24.5%)	43 (18.5%)	49 (34.3%)	

Depressive symptoms were measured with the Patient Health Questionnaire (PHQ-9), range 0–27. Chronotype was measured as the midpoint of sleep on free days corrected for sleep debt on workdays (range, 0–24); higher values indicate greater eveningness, or the tendency to sleep late and wake up late. Characteristics of participants who completed any follow-up and of those who did not were compared with the Kruskal–Wallis, chi-squared, and Fisher's tests

Table 2 Association between social media use at T1 and longitudinal depressive symptoms

Independent variables	Regression coefficient	95% confidence intervals
<i>Initial model</i>		
Time point	0.296	−0.030, 0.623
Age in years	−0.208	−0.464, 0.048
Male sex	−0.523	−2.077, 1.031
Social media use, hours/day	0.525	0.259, 0.791
<i>Adjusted model</i>		
Time point	0.316	−0.010, 0.642
Age in years	−0.238	−0.484, 0.007
Male sex	−1.179	−2.704, 0.347
Social media use, hours/day	0.403	0.142, 0.663
Physical activity	−0.102	−0.395, 0.191
Greenspace exposure, days/week	−0.189	−0.449, 0.072
Cannabis use	2.254	0.880, 3.628
Chronotype	0.723	0.335, 1.111

Depressive symptoms were measured with the Patient Health Questionnaire, range 0–27, at three time points (T1, T2 and T3). All independent variables were measured at T1 ($n=376$). Physical activity was measured in METS-minutes/week then rescaled to 1:1000. Cannabis use was dichotomized as at least once/week vs. none. Chronotype was measured as the midpoint of sleep on free days (range, 0–24); higher values indicate greater eveningness



Estimated marginal means from a linear mixed model ($n=376$) of depressive symptoms at T1, T2 and T3. Depressive symptoms were measured with the Patient Health Questionnaire, range: 0–27, at three time points (T1, T2 and T3). The independent variable is social media use and its interaction with time, after adjusting for age, sex, and other health behaviors.

Fig. 1 Trajectory of depressive symptoms as a function of social media use at T1

of social media use at T1, depressive symptoms were higher by ~0.5 point on the PHQ-9. In male participants, this association was in the opposite direction and not significant: coefficient = -0.234 (95% CI, $-0.821, 0.353$).

Longitudinal Social Media Use as a Function of Depressive Symptoms at T1

More depressive symptoms at T1 were associated with greater social media use on average over T1, T2, and T3 (Table 3). The association remained significant after adjusting for other health-related behaviors at T1. Greater greenspace exposure at T1 was associated with less social media use. Social media use significantly decreased over time. Sex, physical activity, cannabis use, and eveningness chronotype at T1 were not significantly associated with social media use. Interactions of depressive symptoms with time (coefficient = -0.008 [95% CI, $-0.033, 0.018$]) and sex (coefficient = -0.066 [95% CI, $-0.144, 0.012$]) were not significant.

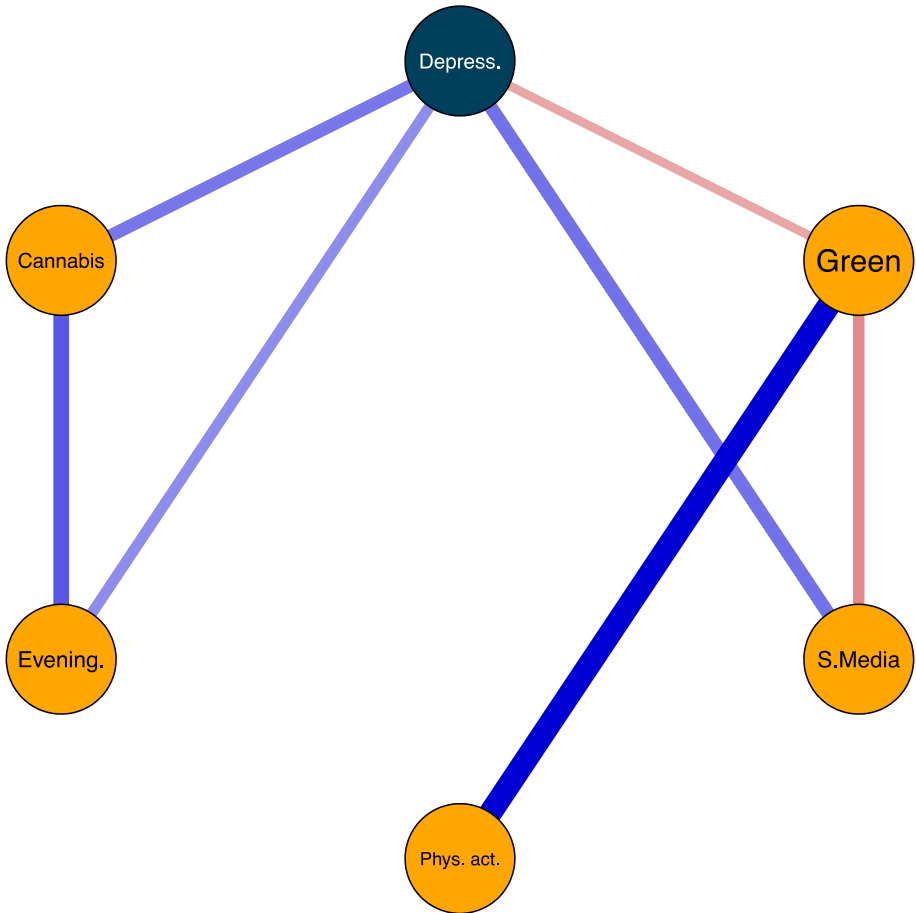
Network Associations

The network model of variables at T1 is presented in Fig. 2. The network identified direct associations of higher levels of social media use with higher levels of depressive symptoms and with lower levels of green space exposure. Other factors associated with higher depression included cannabis use, higher eveningness, and lower greenspace exposure. Physical activity was not directly associated with depression but shared an indirect relationship through its positive association with greenspace exposure. In non-parametric bootstrap analyses, the network edges differed from zero over most iterations (proportion of null estimates ≤ 0.5), supporting their robustness (Table S2; Figures S1-S2).

Table 3 Association between depressive symptoms at T1 and longitudinal social media use

Independent variables	Regression coefficient	95% confidence intervals
<i>Initial model</i>		
Time point	-0.263	-0.413, -0.114
Age in years	-0.072	-0.151, 0.008
Male sex	-0.366	-0.855, 0.123
Depressive symptoms	0.068	0.039, 0.097
<i>Adjusted model</i>		
Time point	-0.264	-0.414, -0.114
Age in years	-0.075	-0.153, 0.004
Male sex	-0.453	-0.949, 0.043
Depressive symptoms	0.052	0.021, 0.082
Physical activity	0.019	-0.076, 0.113
Greenspace exposure, days/week	-0.106	-0.189, -0.023
Cannabis use	0.132	-0.325, 0.589
Chronotype	0.121	-0.006, 0.249

Social media use was measured in hours/day at three time points (T1, T2, and T3). All independent variables were measured at T1 ($n=376$). Physical activity was measured in METS-minutes/week then rescaled to 1:1000. Cannabis use was dichotomized as at least once/week vs. none. Chronotype was measured as the midpoint of sleep on free days (range, 0–24); higher values indicate greater eveningness



Complete data (n=376) at T1. Blue edges indicate positive associations, and red edges indicate negative associations. Thicker edges indicate greater absolute edge weights. Depress. (range 0–27): depressive symptoms. S.Media: social media use (hours/day). Phys. act.: physical activity (MET-minutes/week). Green: frequency of green space exposure (range: 0–7 days/week). Cannabis: frequency of cannabis use (range: 0–7 days/week). Evening.: eveningness chronotype (range: 0–24).

Fig. 2 Network model of cross-sectional associations between depressive symptoms, social media use, and other measures at T1

Network models at T2 and T3 identified fewer associations (Fig. 3). At both T2 and T3, there was an association between greater social media use and more depression, as well as between more physical activity and greater greenspace exposure. At T2, there was also an association between greater social media use and less greenspace exposure (as found at T1). At T3, this association was not present; however, there was an association between greater greenspace exposure and less depression (as found at T1). Other associations were not supported.

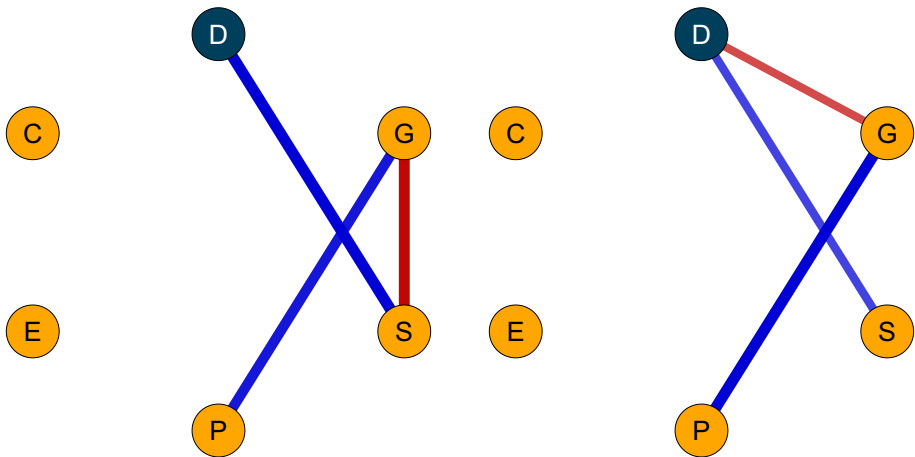


Fig. 3 Network model of cross-sectional associations between depressive symptoms, social media use, and other measures at T2 and T3. Complete data, $n=218$ at T2 (left) and $n=184$, T3 (right). Blue edges indicate positive associations, and red edges indicate negative associations. Thicker edges indicate greater absolute edge weights. Depressive symptoms (range, 0–27); S, social media use (hours/day); P, physical activity (MET-minutes/week); G, frequency of green space exposure (range, 0–7 days/week); C, frequency of cannabis use (range, 0–7 days/week); E, eveningness chronotype (range, 0–24)

Discussion

This study found that greater social media use was associated with higher levels of depression in female young adults, even after controlling for other health-related behaviors at baseline. However, greater use of social media at baseline was not associated with an increase or decrease in depressive symptoms over time. These results show that social media use and depression are associated, supporting the notion that greater social media use may be a marker or correlate of concurrent depressive symptoms, but do not provide evidence that social media use is prospectively related to the course of depressive symptoms. The network analysis identified direct connections of higher levels of social media use with higher levels of depressive symptoms and with lower levels of green space exposure. Cannabis use, higher eveningness, and less greenspace exposure were also associated with more depressive symptoms at T1. These results highlight the interrelation of depression and behavioral and environmental factors in young adults, with potential implications for the early detection of depression and for lifestyle interventions in this population.

The association between greater social media use and higher levels of depression is consistent with previous research. Between-person associations indicate that those individuals with greater digital media use, including social media use, tend to experience more depression. Some studies have also identified within-person associations between changes in digital media use and changes in depression and vice versa, offering stronger evidence for causal effects as participants are compared with themselves, controlling for other individual characteristics, measured or not (Orben et al., 2022). Other studies, however, have found minimal or no association at the within-person level, leading to uncertainty about the true magnitude of effects of digital media use on youth mental health (Houghton et al., 2018; Orben et al., 2019; Steinsbekk et al., 2023). Interventions such as a brief abstinence period, an approach borrowed from substance use therapeutic interventions, have been suggested

to improve mental health outcomes, with some experimental evidence of improvements in well-being, mood, and anxiety after 1 to 2-week periods of abstinence (Hunt et al., 2018; Lambert et al., 2022). However, other research has found no evidence of increased well-being with these “digital detoxes” (Przybylski et al., 2021).

Contrary to our hypothesis, adjusting for other health-related behaviors did not explain the association between social media use and depression, which may be influenced by yet other factors. In some individuals, neurobiological factors may predispose to higher social media use and depression, or mediate the relationship between the two (Becker et al., 2017; Yuan et al., 2017) in youth with high screen exposure. For example, increased positive functional connectivity between the striatum and the frontal executive network has been associated with higher efficiency in inhibitory control during development (Liston et al., 2006; Rubia et al., 2006), which could be a shared correlate of social media use and mental health. Socio-environmental factors, including lower socioeconomic status, racialized minority status, and individual exposure to social adversity may also be shared risk factors for excessive social media use and mental health problems (Nagata et al., 2022).

Other associations in the network pointed to research directions for intervention in young adults. Higher exposure to green spaces was associated with lower levels of depression at T1 and T3. However, in the longitudinal models, greenspace exposure was not significantly associated with average depressive symptoms across T1, T2, and T3. There is extensive research on green spaces and health (Zhang, J. et al., 2020a). Despite relatively less research on mental health specifically, and heterogeneity in the outcome measures used, cumulative evidence points towards an association between positive mental health outcomes and green spaces (Bezold et al., 2018; Engemann et al., 2019; Gascon et al., 2015; Zhang, Y. et al., 2020b). In a randomized-controlled study of 37 adults with major depressive disorder, a single 60-min walk in nature was associated with an improvement in negative affect relative to an urban walk (Watkins-Martin et al., 2022). The lack of consistent associations of greenspace exposure with depression in the present study may stem from various factors, including overall low levels of depression in the sample (compared to clinical populations) and the impact of 3-month intervals between time points, which may be too long to capture momentary effects of health-related behaviors.

Physical activity was not directly associated with longitudinal depressive symptoms, but it was indirectly related to depression via green space exposure in the network models. One concern raised by the World Health Organization (WHO) (World Health Organization. Regional Office for the Eastern Mediterranean, 2020) is that the increase in screen time could replace healthy behaviors such as physical activity and sleep. A population-level prospective cohort study in the UK found that theoretically replacing 1 h of social media use with a team sport was associated with reductions in emotional distress scores (Kandola et al., 2022). National guidelines on physical activity recommend 150 min of moderate intensity or 75 min of vigorous-intensity aerobic physical activity weekly or an equivalent combination, as well as limited sedentary time (Piercy et al., 2018). In the U.S., sedentary behavior significantly increased from 2007 to 2016, coinciding with an increase in smartphone ownership (Du et al., 2019). Sedentary behavior has implications for mental health (Gibson et al., 2017), independent of the amount of physical activity (Bankoski et al., 2011; Maher et al., 2013; Zhai et al., 2015). In the present study, it is likely that those who spent more time doing physical activity or visiting green spaces, and less time on social media, were also less sedentary.

The association of eveningness with depression in the network analysis at T1 is consistent with previous research. A cross-sectional study in a non-clinical sample identified associations between eveningness and depression, mediated by affect and behavioral activation

(Hasler et al., 2010). In the present network model, eveningness was not directly associated with social media use, but it was found elsewhere to be cross-sectionally associated with greater computer use (Patterson et al., 2016). Given the associations of eveningness with depression and cannabis use in the present network results, it is possible that eveningness is only indirectly associated with social media use via these other factors. The lack of association involving eveningness in network models at T2 and T3 may be due to the lower sample size, which constrained statistical power.

This study in part supports the displacement theory, as it shows that greater social media use was associated with less time spent in green spaces. The types of activities being displaced by social media may vary. For example, a study of youth undergoing abstinence found that abstinence from social media increased time engaged in other activities such as working, cooking, or cleaning (Hall et al., 2019). Social media may also be a substitute for other types of screen time (Hall & Liu, 2022). While social media use may displace certain activities, it may also amplify or increase other exposures, such as those related to interpersonal relationships, both positive and negative. As our study did not assess the types and contents of social media use, we could not examine their differential content and engagement effects, which may be more important than the overall time spent on social media (Morningstar et al., 2023).

Overall, the findings support the interrelationship of health-related behaviors, which is important to consider in future research given that various health-related behaviors may be correlates or predisposing factors for social media use and mental health outcomes. However, these findings do not provide evidence that greater social media use predicts increasing depressive symptoms over time. The lack of a longitudinal association between social media use at T1 and changes in depression across time may suggest that time spent on social media is more a marker of than a cause of depression, meaning that those people with a tendency to become depressed may be also more likely to use social media and vice-versa. Other factors may explain this association, such as pre-existing mental health, behavioral, and environmental factors. Importantly, associations between social media use and depression are also likely to vary depending on the types of content and actions that are accessed via social media. Specific aspects of social media use may have positive or negative impacts on depression, and these aspects may manifest to different degrees between individuals and over time. Our findings highlight the importance of taking an integrative approach when examining social media use and health-related behaviors, with the promotion of healthy behaviors as a potential intervention to be investigated in future research to address the various factors that could impact depression in tandem with social media use.

Limitations of this study include convenience sampling and the reliance on self-reported data. More “objective” measures of health-related behaviors may be obtained with digital phenotyping, a method of data collection that can include logs of social media use, sleep indices, and geolocation, via the use of smartphones or wearables. Importantly, we lacked data on the different types and contents of social media that were used, which may be differentially associated with depressive symptoms and other health-related behaviors. Lastly, the present sample was relatively small, which limited statistical power and the reliability of network estimates. Due to online convenience sampling, the sample included a majority of female individuals, further limiting the generalizability of the results to the general population. Attrition, which was predicted by male sex, cannabis use, and more depressive symptoms at baseline, may have influenced the regression (if data was missing not at random) and network estimates.

In conclusion, this study found concurrent associations of higher social media use with higher depression scores, as well as with other health-related behaviors including less time

spent in green spaces. Despite their interrelationship, health-related behaviors may not be sufficient to explain the association between social media use and depression, which may instead reflect the influence of other individual, media-related, and socio-environmental factors.

Supplementary Information The online version contains supplementary material available at <https://doi.org/10.1007/s11469-024-01313-0>.

Funding This study received funding from the Stratas Foundation awarded to Dr. Paquin. Dr. Vidal receives funding from the NIDA/AACAP Physician Scientist Training Program in Substance Abuse Research, supported by NIDA Career Development Award (K12DA000357). Dr. Geoffroy holds a Canada Research Chair (Tier 2) in Youth Mental Health and Suicide Prevention. Dr. Paquin is supported by an award from the Quebec Network on Suicide, Mood Disorders and Associated Disorders.

Data Availability The data that support the findings of this study are available from the last author, Vincent Paquin (vincent.paquin2@mail.mcgill.ca), upon reasonable request.

Declarations

Ethics Approval and Consent to Participate All procedures followed were in accordance with the ethical standards of the responsible committee on human experimentation (institutional and national) and with the Helsinki Declaration of 1975, as revised in 2000. Informed consent was obtained from all participants for being included in the study.

Conflict of Interest The authors declare no competing interests.

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